

# Probability Lecture Student Notes

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## Asynchronous Lecture

### Expectations of Sums

Recall that:

- In the discrete case,  $E[X] = \sum_i x_i p(x_i)$ ,
- In the continuous case,  $E[X] = \int x f(x) dx$

If random variables  $X, Y$  have joint pdf  $f(x, y)$ , then

$$E[g(X, Y)] = \int \int g(x, y) f(x, y) dx dy$$

$g(X, Y)$  may be anything from  $X$ , to  $Y$ , to  $2XY^2$ .

**Example:**  $g(x, y) = x + y$

$$\begin{aligned} E[g(X, Y)] &= \int \int (x + y) f(x, y) dx dy \\ &= \int \int x f(x, y) dx dy + \int \int y f(x, y) dx dy \\ &= \int x f_x(x) dx + \int y f_y(y) dy \\ &= E[X] + E[Y] \end{aligned}$$

Linearity of expectation  $\Rightarrow E[X_1 + X_2 + \dots + X_n] = E[X_1] + E[X_2] + \dots + E[X_n]$

*Note:* this is not necessarily the case for infinite sums. Only if the following exchange of  $\lim_{n \rightarrow \infty}$  and  $\sum_{i=1}^n x_i$  holds is it true.

$$\begin{aligned} E\left[\sum_{i=1}^{\infty} x_i\right] &= E\left[\lim_{n \rightarrow \infty} \sum_{i=1}^n x_i\right] \\ &= \lim_{n \rightarrow \infty} \sum_{i=1}^n x_i E[x_i] \end{aligned}$$

Guaranteed to be exchangeable if  $\forall i, x_i \geq 0$ ; or if  $\sum_{i=1}^n E[|x_i|] < \infty$  (absolutely convergent)

## Covariance and Variance of Sums

**Definition: Covariance**

$$\begin{aligned} Cov(X, Y) &= E[(X - \mu_x)(Y - \mu_y)] \\ &= E[XY] - E[X]E[Y] \\ &= \int \int xyf(x, y) dx dy - E[X]E[Y] \end{aligned}$$

If  $X$  and  $Y$  are independent, then  $Cov(X, Y) = 0$

*Note:* this is not necessarily true in reverse!

## Covariance Properties

1.  $Cov(X, Y) = Cov(Y, X)$
2.  $Cov(X, X) = Var(X)$
3.  $Cov(aX, Y) = aCov(X, Y)$
4.  $Cov(\sum_{i=1}^n x_i, \sum_{j=1}^m y_j) = \sum_{i=1}^n \sum_{j=1}^m Cov(x_i, y_j)$

If we combine properties 4 and 2,

$$\begin{aligned} Var\left(\sum_{i=1}^n x_i\right) &= Cov\left(\sum_{i=1}^n x_i, \sum_{i=1}^n x_i\right) \\ &= \sum_{i=1}^n \sum_{j=1}^n Cov(x_i, x_j) \end{aligned}$$

We can pull out where  $x_i = x_j$ , which occurs when  $i = j$

$$= \sum_{i=1}^n Var(x_i) + \sum_{i \neq j} Cov(x_i, x_j)$$

And then add property 1,

$$= \sum_{i=1}^n Var(x_i) + 2 \sum_{i < j} Cov(x_i, x_j)$$

If the  $x_i$  are independent,  $Var(\sum_{i=1}^n x_i) = \sum_{i=1}^n Var(x_i)$

## Correlation

**Definition: Correlation**

$$\rho(x, y) = \frac{Cov(X, Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}$$

We can compare this to linear algebra. Recall that the dot (inner) product  $\vec{x} \cdot \vec{y} = \sum_{i=1}^n x_i y_i$ , and that the norm  $|\vec{x}| = \sqrt{\vec{x} \cdot \vec{x}}$ . We can also rewrite the dot product as

$$\vec{x} \cdot \vec{y} = |\vec{x}| |\vec{y}| \cos \theta$$

and replace each term with parts from the definition of covariance, such that

$$\begin{array}{ll} \vec{x} \cdot \vec{y} \rightarrow Cov(X, Y) & |\vec{x}| \rightarrow std(x) \\ |\vec{x}|^2 \rightarrow var(x) & \cos \theta \rightarrow \rho(x, y) \end{array}$$

Then, we end up with

$$\cos \theta = \frac{\vec{x} \cdot \vec{y}}{|\vec{x}| |\vec{y}|}$$

If  $\rho(x, y) = 0$ , we say that  $x$  and  $y$  are *uncorrelated*,  
 or in vector terms, they are *orthogonal* ( $\cos \theta = 0 \Rightarrow \theta = \frac{\pi}{2}$ )  
 If  $\rho(x, y) = 1$ , we know that  $x = aX + b$ ,  
 or in vector terms, they are *colinear* ( $\cos \theta = 1 \Rightarrow \theta = 0$ )

## Synchronous Lecture

### Expectations of Sums

Recall that for a collection of Random Variables  $X_1 + X_2 + \dots + X_n$ , if  $Y = X_1 + X_2 + \dots + X_n$ , then  $E[X] = E[X_1 + X_2 + \dots + X_n] = \sum E[X_i]$

**Definition: *Sample mean***

$$\bar{x} = \frac{1}{n}(x_1 + x_2 + \dots + x_n)$$

If  $X_i$  are IID Random Variables, with each  $E[x_i] = \mu_i$ ,

$$\begin{aligned} E[\bar{x}] &= \frac{1}{n}(\sum E[X_i]) \\ &= \frac{1}{n} \cdot \mu \\ &= \mu \end{aligned}$$

note: in statistics, by observing data, we infer property of the R.V.process.  
 For here, we estimate the mean of those Random Variables by taking average of data.

$$\begin{aligned} E[g(X, Y)] &= \int \int (x + y)f(x, y)dxdy \\ &= \int \int xf(x, y)dxdy + \int \int yf(x, y)dxdy \\ &= \int xf_x(x)dx + \int yf_y(y)dy \\ &= E[X] + E[Y] \end{aligned}$$

## Covariance

**Definition: Covariance**

$$Cov(X, Y) = E[(X - \mu_x)(Y - \mu_y)]$$

If  $X, Y$  independent, then  $Cov(X, Y) = 0$

## Correlation

**Definition: Correlation**

$$\rho(x, y) = \frac{Cov(X, Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}$$

with  $\rho(x, y) \in [-1, 1]$ .

If we consider  $Cov(aX, Y)$ ,

$$\begin{aligned} Cov(aX, Y) &= aCov(X, Y) \\ \rho(aX, Y) &= \frac{aCov(X, Y)}{\sqrt{var(aX)}\sqrt{var(Y)}} \\ &= \frac{aCov(X, Y)}{\sqrt{a^2 var(X)}\sqrt{var(Y)}} \\ &= \frac{a}{|a|} \rho(X, Y) \end{aligned}$$

note: correlation shows that we scale the random variables by some characteristic size, and then we compute Covariances.

## Sample Variance

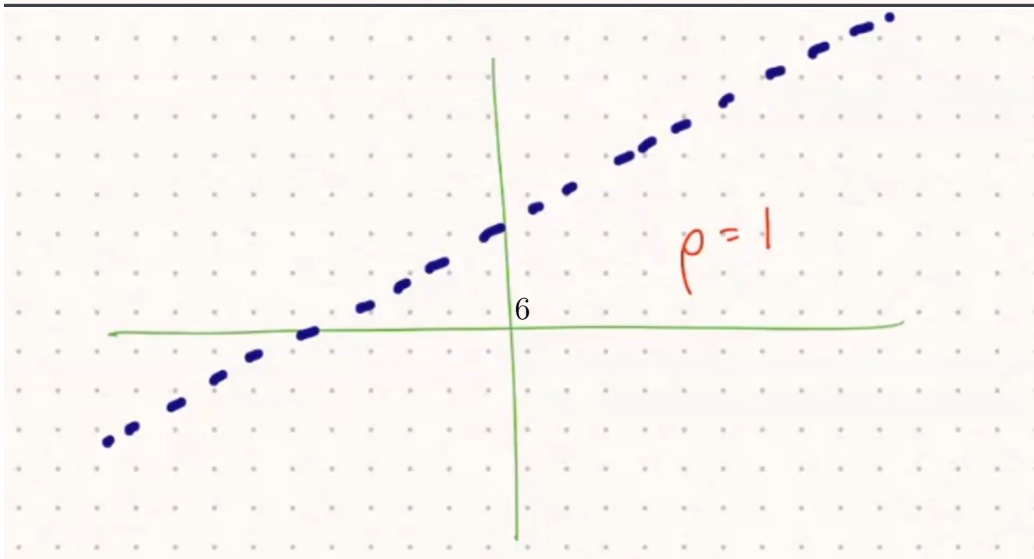
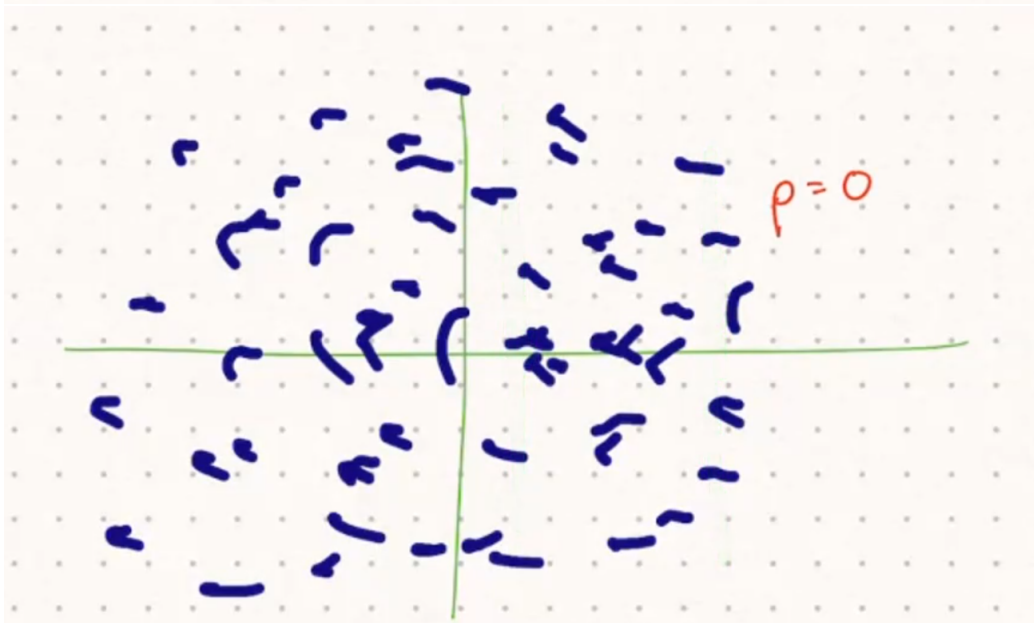
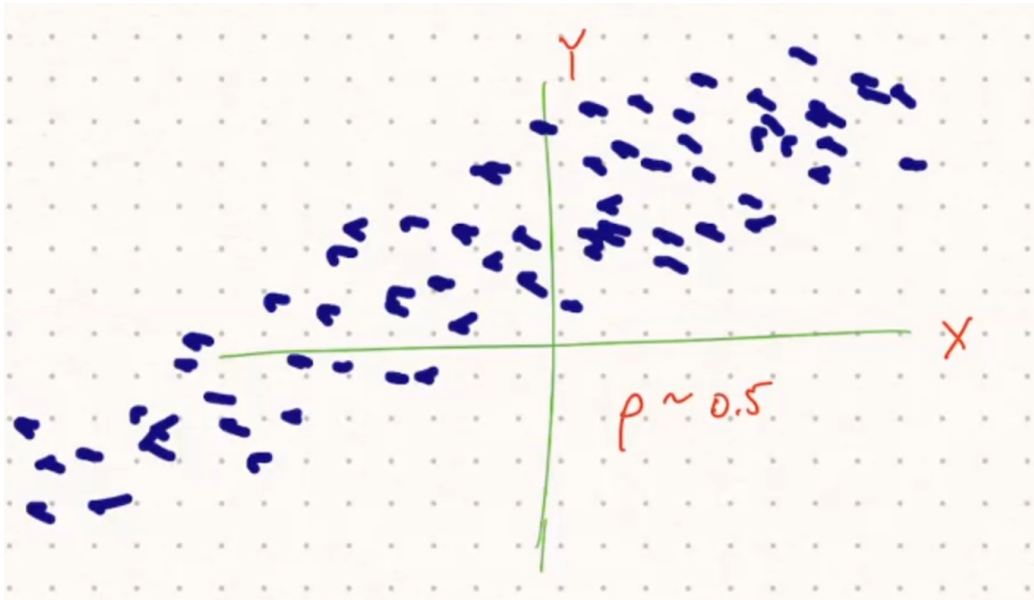
**Definition: Sample Variance**

If  $X_1 + X_2 + \dots + X_n$  are IID Random Variables with mean  $\mu$  and variance  $\sigma^2$ , then the sample variance is given by

$$S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n - 1}$$

Note that  $E[S^2] = \sigma^2$ .

Student question about how to visualize these computations. There are three graphs of visualization of different covariances.



## Exercises

### Theoretical exercise 7.23

If  $Y = a + bX$ , what is  $\rho(X, Y)$ ?

$$\begin{aligned} \text{Var}(X) &= \sigma^2 \\ \text{Var}(Y) &= \text{Var}(a + bX) \\ &= b^2\sigma^2 \end{aligned}$$

$$\begin{aligned} \text{Cov}(X, Y) &= \text{Cov}(X, a + bX) \\ &= E(X - \mu)[a + bX - (a + b\mu)] \\ &= E[(X - \mu)(bX - b\mu)] \\ &= bE[(X - \mu)(X - \mu)] \\ &= b\sigma^2 \end{aligned}$$

$$\begin{aligned} \rho(X, Y) &= \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}} \\ &= \frac{b\sigma^2}{\sigma \cdot |b|\sigma} \\ &= \frac{b}{|b|} \end{aligned}$$

1 if  $b > 0$ , and -1 if  $b < 0$

### Theoretical exercise 7.4

Let  $X$  be a Random Variable with

$$\begin{aligned} E[X] &= \mu < \infty \\ \text{Var}[X] &= \sigma^2 < \infty \end{aligned}$$

and  $g = g(x)$  is twice differentiable, how to approximate  $E[g(x)]$  ?

By Taylor Series, we know that

$$\begin{aligned}g(x) &\approx g(\mu) + g'(\mu)(x - \mu) + \frac{g''(\mu)}{2}(x - \mu)^2 \\E[g(x)] &\approx E[g(\mu) + g'(\mu)(x - \mu) + \frac{g''(\mu)}{2}(x - \mu)^2] \\&= g(\mu) + g'(\mu)E[X - \mu] + \frac{g''(\mu)}{2}E[(X - \mu)^2] \\&= g(\mu) + \frac{g''(\mu)}{2}\sigma^2\end{aligned}$$

$$\underline{E[g(x)] \approx g(\mu) + \frac{g''(\mu)}{2}\sigma^2}$$