

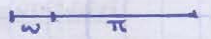
PCPP (Assignment Testers)

Let  $P \subseteq \{-1, 1\}^m$  be some property of  $m$ -bit strings (e.g., the graph represented by the string is 3-colorable). Assume someone tries to convince us that a string  $w$  has the property  $P$ . However, we can only look at a constant # of bits.

To help us the prover has a "proof" that  $w \in P$ .

Def: A property  $P$  of  $m$ -bit strings has PCPPs of length  $l(m)$  if there exists a tester  $T$  making  $O(1)$  non-adaptive queries to a string " $w, \pi$ " where  $\pi \in \{-1, 1\}^{l(m)}$  s.t.

(1). If  $w \in P$  then  $\exists \pi_w$  s.t.  $\Pr[T \text{ accepts } w, \pi_w] = 1$ .



(2). If  $w$  is  $\epsilon$ -far from  $P$ , then  $\forall \pi$ ,  $\Pr[T \text{ accepts } w, \pi] < 1 - \Omega(\epsilon)$ .

Thm: [BHSV+DR] Every property  $P \subseteq \{-1, 1\}^m$  has PCPPs of length  $2^m$  with only 3 queries.

Remark: This is about 25% of Dinur's proof of the PCP thm.

Proof: We will ~~expect~~ identify  $\{-1, 1\}^m$  with  $[n] = \{1, \dots, n\}$  where  $n = 2^m$ . So  $w \in [n]$  and  $P \subseteq [n]$ . We set  $\pi_w$  to be the truth table of  $\chi_{\{w\}}: \{0, 1\}^m \rightarrow \{-1, 1\}$ .

The tester should check: (1).  $\pi$  is a dictator on a coordinate in  $P$ . (2). This coordinate is  $w$ . We achieve this by running one of the following tests with equal probability: (1). The "subset of dictators" test using  $P$  on  $\pi$ .

(2). Choose  $i \in [m]$  uniformly. Use local decoding on  $\pi$  to compute its value on the string  $x \in \{0, 1\}^m$  given by  $x_j = 1$  iff  $j_i = -1$  (where we think of  $j$  as both  $j \in [n]$  and  $j \in \{-1, 1\}^m$ ). Also query  $w_i$  and accept if both are equal.

Hardness of Approximation

Influence: For a Boolean function  $f: \{0, 1\}^m \rightarrow \{-1, 1\}$  we define the influence of the  $i$ th coordinate to be  $\text{Inf}_i(f) = \Pr_x [f(x) \neq f(x \oplus e_i)]$ .

Examples: 1. Dictatorship  $\chi_{\{i\}}: \text{Inf}_i = 1$ ;  $\text{Inf}_j = 0$  (for  $j \neq i$ ).

2. Parity:  $\forall i, \text{Inf}_i = 1$ . 3. Majority:  $\forall i, \text{Inf}_i \approx \frac{1}{\sqrt{m}}$ .



4. Constant function:  $\forall i, \text{Inf}_i = 0$ .

Let's try to express  $\text{Inf}$  in terms of the Fourier coeff. Define  $g(x) = f(x \oplus e_i)$ . Then, as we saw,  $\hat{g}(s) = \hat{f}(s)$  if  $i \notin s$ , and  $\hat{g}(s) = -\hat{f}(s)$  if  $i \in s$ . Now, define  $h(x) = f(x) - g(x)$ .

Then  $h$  is 0 if  $f(x) = f(x \oplus e_i)$  and  $\pm 2$  otherwise. Hence,  $\text{Inf}_i(f) = \frac{1}{4} E[h(x)^2] =$

$$= \frac{1}{4} \sum_s \hat{h}(s)^2 = \frac{1}{4} \sum_s (\hat{f}(s) - \hat{g}(s))^2 = \sum_{i \in s} \hat{f}(s)^2.$$



Def: For  $f: \{0,1\}^n \rightarrow \mathbb{R}$  define  $\text{Inf}_i(f) = \sum_{s \neq i} \hat{f}(s)^2$ .

Total influence: For  $f: \{0,1\}^n \rightarrow \mathbb{R}$  define its total influence to be  $I(f) = \sum_i \text{Inf}_i(f)$ .

Claim:  $I(f) = \sum_s |s| \cdot \hat{f}(s)^2$

Examples: Dictator - 1, Parity -  $n$ , Maj -  $\sqrt{n}$ , Const - 0.

Claim: For any  $f: \{0,1\}^n \rightarrow [-1,1]$  with  $E[f] = 0$ ,  $I(f) \geq 1$ .

Proof:  $I(f) = \sum_s |s| \cdot \hat{f}(s)^2 \geq \sum_s \hat{f}(s)^2 = 1$ .  $\square$

Ex: Moreover, if  $I(f) = 1$  and  $E[f] = 0$  then  $f$  is a dictator or antidictator.

Attenuated influence: Consider a situation where with some prob.  $\delta > 0$  each vote is miscounted. In such a situation the expected outcome is  $T_{1-2\delta} f(x)$  (and not  $f(x)$  as before). So here we should look at  $\frac{1}{4} E_x [(T_{1-2\delta} f(x) - T_{1-2\delta} f(x \oplus e_i))^2] = \sum_{s \neq i} (1-2\delta)^{|s|} \cdot \hat{f}(s)^2$ .

Def: For  $-1 \leq p \leq 1$  define the  $p$ -attenuated influence as  $\text{Inf}_i^{(p)}(f) = \sum_{s \neq i} p^{|s|} \cdot \hat{f}(s)^2$ .

Examples:  $\chi_{\{i\}}$ :  $\text{Inf}_i^{(p)} = \begin{cases} p & j=i \\ 0 & \text{o.w.} \end{cases}$ ;  $\chi_{[n]}$ :  $\forall i, \text{Inf}_i^{(p)} = p^n$  (very small).

Maj:  $\forall i, p$  constant,  $\text{Inf}_i^{(p)} \approx \frac{1}{n}$ ;  $\text{Inf}_i^{(p)}(f)$  is increasing as  $p$  goes from 0 to 1.

$\text{Inf}_i^{(1)}(f) = \text{Inf}_i(f)$ ;  $\text{Inf}_i^{(0)}(f) = \hat{f}(i)^2$ .

Proposition: For  $f: \{0,1\}^n \rightarrow [-1,1]$ ,  $\sum_{i=1}^n \text{Inf}_i^{(p)}(f) \leq 1/\delta$ , and hence  $\#\{i: \text{Inf}_i^{(1-\delta)}(f) \geq \epsilon\} \leq 1/\epsilon\delta$ .  
( $0 < \delta < 1$ )

Proof: First notice that  $\forall k \geq 0, k \cdot (1-\delta)^k \leq \sum_{i=0}^{\infty} (1-\delta)^i = 1/\delta$ .

So,  $\sum_{i=1}^n \text{Inf}_i^{(1-\delta)}(f) = \sum_{i=1}^n \sum_{s \neq i} (1-\delta)^{|s|} \cdot \hat{f}(s)^2 = \sum_s |s| \cdot (1-\delta)^{|s|} \cdot \hat{f}(s)^2 \leq \frac{1}{\delta} \cdot \sum_s \hat{f}(s)^2 = \frac{1}{\delta} \cdot E[f(x)^2] \leq \frac{1}{\delta}$ .  $\square$

Quasirandomness: We say that a function  $f: \{0,1\}^n \rightarrow [-1,1]$  is  $(\epsilon, \delta)$ -quasirandom if  $\forall i \in [n], \text{Inf}_i^{(1-\delta)} < \epsilon$ .

Examples: Maj, Parity, Random are all quasirandom. Dictator,  $5$ -junta are not.

Thm: If the Hast-Odds accepts with prob.  $> \frac{1}{2} + \frac{1}{2}\epsilon$  the function  $h: \{0,1\}^n \rightarrow [-1,1]$ , then  $h$  is not  $(\epsilon^2, \delta)$ -quasirandom.

Proof: As we saw the acceptance prob of Hast-Odds is  $\frac{1}{2} + \frac{1}{2} \sum_{|s| \text{ odd}} (1-2\delta)^{|s|} \cdot \hat{h}(s)^2$ .

Therefore,  $\epsilon < \sum_{|s| \text{ odd}} (1-2\delta)^{|s|} \hat{h}(s)^2 \leq \max_{|s| \text{ odd}} \{(1-2\delta)^{|s|} \hat{h}(s)\} \cdot \sum_{|s| \text{ odd}} \hat{h}(s)^2 \leq \max_{|s| \text{ odd}} \{(1-2\delta)^{|s|} \hat{h}(s)\}$ .

Therefore  $\exists S^* \neq \emptyset$ , for which  $(1-2\delta)^{|S^*|} \cdot \hat{h}(S^*) > \epsilon$ .

Let  $i \in S^*$  be any index in  $S^*$ . Then,  $\text{Inf}_i^{(1-\delta)}(h) = \sum_{s \neq i} (1-\delta)^{|s|} \hat{h}(s)^2 \geq (1-\delta)^{|S^*|} \cdot \hat{h}(S^*)^2 > \epsilon^2$ .  $\square$

We are given a list of equations mod 2 with 3 variables in each.

Assume we are promised that there's a solution satisfying 99% of

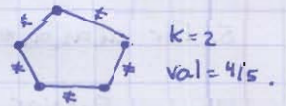
equations. Can we find it?

$$\begin{cases} x_3 + x_9 + x_{10} = 0 \pmod{2} \\ x_2 + x_6 + x_7 = 1 \pmod{2} \\ \vdots \\ x_5 + x_{10} + x_4 = 0 \pmod{2} \end{cases}$$



We can satisfy 50% by assigning it randomly. Can we do 51%? No, it is NP-hard. [Hastad 97]

Label Cover Problem: The label cover problem - we are



given a graph and a constraint  $\phi_e: [k] \times [k] \rightarrow \{F, T\}$

associated to each  $e \in E$ . Our goal is to assign a value in  $[k]$  to each vertex so as to maximize the number of satisfied constraints.

Remark: This is an instance of 2-CSP (Constraint Satisfaction Problem).

Thm (PCP Theorem [AS92, ALMSS92]) + (Parallel Repetition [Raz95])

$\forall \epsilon > 0, \exists k, s.t.$  it is NP-hard given a label cover with alphabet  $[k]$  to tell if  $val = 1$  or  $val \leq \epsilon$ . Moreover, the graph is bipartite and all constraints are "projections".

Proof outline: The PCP theorem says that it is NP-hard to tell if a given 3CNF is satisfiable or  $< 99\%$  satisfied. We convert a 3CNF formula to label cover:

$$(x_2 \vee \bar{x}_5 \vee x_4) \wedge (x_1 \vee \bar{x}_2 \vee \bar{x}_3) \wedge \dots$$

If the 3CNF is satis. then  $val = 1$ .

If the 3CNF is  $1-\delta$  satis. (for  $\delta = \frac{1}{100}$ )



then  $val \leq 1 - \delta^3$ . We now apply the parallel repetition theorem, which makes  $(1, 1 - \delta^3)$  into  $(1, \epsilon)$ . It works by replacing the label cover instance by a power of itself.

Def: A constraint  $\phi: [k] \times [k] \rightarrow \{T, F\}$  is called unique if for  $\forall a \in [k]$  there exists exactly one  $b \in [k]$  s.t.  $\phi(a, b) = T$  and vice versa. In other words,  $\phi$  is a permutation on  $[k]$ .

Claim: There is a poly-time algorithm to check if  $val = 1$  or not for any given unique-label-cover.

(Note that without "uniqueness" this is NP-hard).

Unique Games Conjecture: [Khot 02]  $\forall \epsilon, \delta > 0, \exists k$  s.t. it is NP-hard given a unique label cover instance, to tell if  $val \geq 1 - \epsilon$  or  $val \leq \delta$ .

### Homework 1 - Solutions

2h.  $g(x_1, \dots, x_n) = f(x_1, x_1, x_2, x_2, \dots, x_n, x_n)$ . The transformation  $f \mapsto g$  is linear (from  $\{0, 1\}^n \rightarrow \mathbb{R}$  to  $\{0, 1\}^{2n} \rightarrow \mathbb{R}$ ). Therefore it suffices to analyze it on  $\{x_s\}$ .  $\chi_\phi \mapsto \chi_\phi$ ;  $\chi_{\{1,1\}} \mapsto \chi_{\{1,1\}}$

$\chi_{\{2,2\}} \mapsto \chi_{\{1,1\}}$ ;  $\chi_{\{1,2\}} \mapsto \chi_\phi$ . In general,  $\chi_s \mapsto \chi_{\pi_i(s)}$  where  $i \in \pi_i(s)$  iff exactly one of  $\{2i-1, 2i\}$  is in  $s$ . So,  $g = \sum_s \hat{f}(s) \chi_{\pi_i(s)}$ .  $\hat{g}(T) = \sum_{s: \pi_i(s)=T} \hat{f}(s)$ .

2d.  $\hat{g}(s) = \hat{f}(s+y)$   $(-1)^{\langle x, y \rangle} \cdot \chi_s(x) = \chi_{s+y}(x)$ .

2j.  $S = \phi$ . In this case,  $g(x) = \mathbb{E}_{y \in \{0,1\}^n} [f(x, y)]$ .  $\chi_T \mapsto \begin{cases} \chi_T & T \in I \\ 0 & \text{o.w.} \end{cases}$  Therefore,  $g(x) = \sum_{T \in I} \hat{f}(T) \chi_T$ .



In general case,  $g(x) = \sum_{T \in \mathbb{F}} \hat{f}(TUS) \chi_T$ .

5. For  $a_1, a_2, a_3 \in \{-1, 1\}$ ,  $NAE(a_1, a_2, a_3) = \frac{3}{4} - \frac{1}{4}a_1a_2 - \frac{1}{4}a_2a_3 - \frac{1}{4}a_1a_3$ .

Hence,  $\Pr[NAE \text{ accepts } f] = \frac{3}{4} - \frac{1}{4}E[f(x)f(y)] - \frac{1}{4}E[f(y)f(z)] - \frac{1}{4}E[f(x)f(z)]$  where for each  $i$ ,

$(x_i, y_i, z_i)$  is chosen uniformly from  $\{001, 010, 100, 011, 101, 110\}$ . By symmetry,

$$= \frac{3}{4} - \frac{3}{4}E[f(x)f(y)] \text{ where for each } i \text{ } (x_i, y_i) \text{ is chosen according to } \begin{cases} 00 & 1/6 \\ 01 & 1/3 \\ 10 & 1/3 \\ 11 & 1/6 \end{cases}$$

Equivalently, we choose  $x$  uniformly and choose  $y = x + w$  where  $w \sim \mu_{1/3}$ .  $\Rightarrow = \frac{3}{4} - \frac{3}{4} \langle f, T_{1/3} f \rangle$ .

$$= \frac{3}{4} - \frac{3}{4} \sum_s \left(-\frac{1}{3}\right)^{|s|} \hat{f}(s)^2$$

3a. Estimate  $\hat{f}(s)$  to within  $\pm \epsilon$  with conf. of  $1 - \delta$ .  $\hat{f}(s) = E[f(x) \chi_s(x)]$ . By Chernoff,

number of samples is  $O(\log 1/\delta / \epsilon^2)$ .

b.  $\sum_{s \in \mathbb{F}^n} \hat{f}(s)^2 = E_x[f(x)^2]$ , so again Chernoff.

$$7. I(f) = \sum_s |s| \hat{f}(s)^2, \quad 4 \cdot \Pr[f=1] \cdot \Pr[f=-1] = 4 \cdot \frac{1}{2}(1 + \hat{f}(\emptyset)) \cdot \frac{1}{2}(1 - \hat{f}(\emptyset)) = 1 - \hat{f}(\emptyset)^2$$

So the inequality is equiv. to  $\hat{f}(\emptyset)^2 + \sum_s |s| \hat{f}(s)^2 \geq 1$ .  
 $\geq \sum_s \hat{f}(s)^2 = 1$

If we color the Hamming cube with two colors in a balanced way, then # of non-monochromatic

edges is  $\geq 2^{n-1}$ . In general  $E(V', V-V') \geq \frac{4|V| \cdot |V-V'|}{2 \cdot 2^n}$ .